

# A look into the factor model black box

Publication lags and the role of hard and soft data in forecasting GDP

Marta Bańbura and Gerhard Rünstler

[mbanbura@ulb.ac.be](mailto:mbanbura@ulb.ac.be)  
Université Libre de Bruxelles

[gerhard.ruenstler@ecb.int](mailto:gerhard.ruenstler@ecb.int)  
Directorate General Research  
European Central Bank

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# Introduction

Forecasting GDP from monthly data: large unbalanced data sets

Dynamic factor models (DFMs) have become a common tool

One drawback is the black box nature of DFMs

- Which series drive the forecast?
- How to handle unbalancedness?
- What are the sources of forecast revisions?

# This paper

We use a DFM by Doz et al. (2005)

Statistics to assess role of individual series

- Kalman filter weights
- Uncertainty measures

Hard & soft data in forecasting euro area GDP

Important role for surveys once unbalancedness is accounted for

$$\begin{aligned}
 x_t &= \Lambda f_t + \xi_t, & \xi_t &\sim \mathbf{N}(0, \Sigma_\xi), \\
 f_t &= \sum_{i=1}^p A_i f_{t-i} + \zeta_t, \\
 \zeta_t &= B \eta_t, & \eta_t &\sim \mathbf{N}(0, I_q).
 \end{aligned}$$

Forecast for monthly GDP  $\hat{y}_t$

In the 3<sup>rd</sup> month of each quarter, evaluate the forecast for quarterly GDP growth,  $\hat{y}_t^Q$

$$\begin{aligned}
 \hat{y}_t &= \beta' f_t \\
 \hat{y}_t^Q &= \frac{1}{3}(\hat{y}_t + \hat{y}_{t-1} + \hat{y}_{t-2})
 \end{aligned}$$

$$\begin{bmatrix} x_t \\ y_t^Q \end{bmatrix} = \begin{bmatrix} \Lambda & 0 & 0 \\ 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} f_t \\ \hat{y}_t \\ \hat{y}_t^Q \end{bmatrix} + \begin{bmatrix} \xi_t \\ \varepsilon_t^Q \end{bmatrix}$$

$$\begin{bmatrix} I_r & 0 & 0 \\ -\beta' & 1 & 0 \\ 0 & -\frac{1}{3} & 1 \end{bmatrix} \begin{bmatrix} f_{t+1} \\ \hat{y}_{t+1} \\ \hat{y}_{t+1}^Q \end{bmatrix} = \begin{bmatrix} A_1 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & \Xi_{t+1} \end{bmatrix} \begin{bmatrix} f_t \\ \hat{y}_t \\ \hat{y}_t^Q \end{bmatrix} + \begin{bmatrix} \zeta_{t+1} \\ 0 \\ 0 \end{bmatrix}$$

# Kalman filter & smoother

For state space form

$$\begin{aligned}z_t &= W\alpha_t + u_t, & u_t &\sim N(0, \Sigma_u) \\ \alpha_{t+1} &= T_t\alpha_t + v_t, & v_t &\sim N(0, \Sigma_v),\end{aligned}$$

and any data  $\mathcal{Z}_t^{-j}$  the KF provides MMSE estimates of  $\alpha_{t+h}$ ,

$$\begin{aligned}a_{t+h|t}^{-j} &= \mathbb{E} \left[ \alpha_{t+h} | \mathcal{Z}_t^{-j} \right] \\ P_{t+h|t}^{-j} &= \mathbb{E} \left[ a_{t+h|t}^{-j} - \alpha_{t+h} \right] \left[ a_{t+h|t}^{-j} - \alpha_{t+h} \right]',\end{aligned}$$

Flexible handling of missing observations

## Euro area data set

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Real activity	32	
Industrial production		6 weeks
Retail sales		6 weeks
Labour market		6-8 weeks
Surveys (EC)	22	0 weeks
Business		
Consumer		
Retail & construction		
Financial data	22	0 weeks
Exchange & interest rates		
Stock price indices		
Other		

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**Recursive pseudo real-time forecasts**  
(End-of-month)

	Example	Real data	Surveys	Financial
	Q2			
Q(-2) M1	(Oct)			
Q(-2) M2	(Nov)			
Q(-2) M3	(Dec)			
Q(-1) M1	Jan			
Q(-1) M2	Feb			
Q(-1) M3	Mar			
Q(0) M1	Apr			
Q(0) M2	May			
Q(0) M3	Jun			
<b>Q(+1) M1</b>	<b>Jul</b>			
Q(+2) M2	Aug			



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Q(0) M2	May			
Q(0) M3	Jun			
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Q(+2) M2	Aug			

# Forecast weights

Express forecasts  $\hat{y}_{t+h|t}^Q$  as the weighted sum of available observations in  $\mathcal{Z}_t$

Algorithm by Harvey and Koopman (2003) to calculate weights

$$\hat{y}_{t+h|t}^Q = \sum_{k=0}^{t-1} \omega_{k,t}(h) z_{t-k} ,$$

Weights are time-invariant for our definition of  $\mathcal{Z}_t$ .

Inspect

- Cumulative forecast weights  $\sum_{k=0}^{t-1} \omega_{k,i}(h)$  for series  $i$ ,
- Historical contributions of series  $i$  to the forecast

# Uncertainty measures

Define *subsets of indicators*  $x_t = (x_t^{1'}, x_t^{2'}, x_t^{3'})'$

$x_t^{-j}$  : all observations of  $x_t^j$  eliminated

$$\mathcal{Z}_t^{-j} = \left\{ x_s^{-j} \right\}_{s=1}^t$$

Marginal contribution to forecast precision of  $x_t^{-j}$  :

difference in precision from data  $\mathcal{Z}_t$  and  $\mathcal{Z}_t^{-j}$

*Advantage: no re-estimation (maintain original factor loadings)!*

**Recursive pseudo real-time forecasts**  
(End-of-month)

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Q(-2) M2				
Q(-2) M3				
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Q(0) M1	Apr			
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Q(0) M3	Jun			
Q(+1) M1	Jul			
Q(+2) M2	Aug			

# Filter uncertainty (Giannone et al., 2005)

Decompose the variance of the forecast error for  $y_{t+h}^Q$

$$\text{var}(\hat{y}_{t+h|t}^{Q,-j} - y_{t+h}^Q) = \pi_{t+h|t}^{-j} + \sigma_{\varepsilon}^2 ,$$

- $\sigma_{\varepsilon}^2$  is residual uncertainty
- $\pi_{t+h|t}^{-j}$  is the uncertainty from  $f_{t+h|t}^{-j}$ .

Our SSF:  $\pi_{t+h|t}^{-j}$  is obtained from  $P_{t+h|t}^{-j}$ .

Inspect increase in  $\pi_{t+h|t}^{-j}$  against  $\pi_{t+h|t}$ .

# RMSE

Filter uncertainty measure ignores parameter uncertainty

Consider out-of-sample forecasts based on recursive parameter estimates

But: same parameters for all data  $\mathcal{Z}_t^{-j}$ .





Chart 1: Cumulative forecast weights across data sets

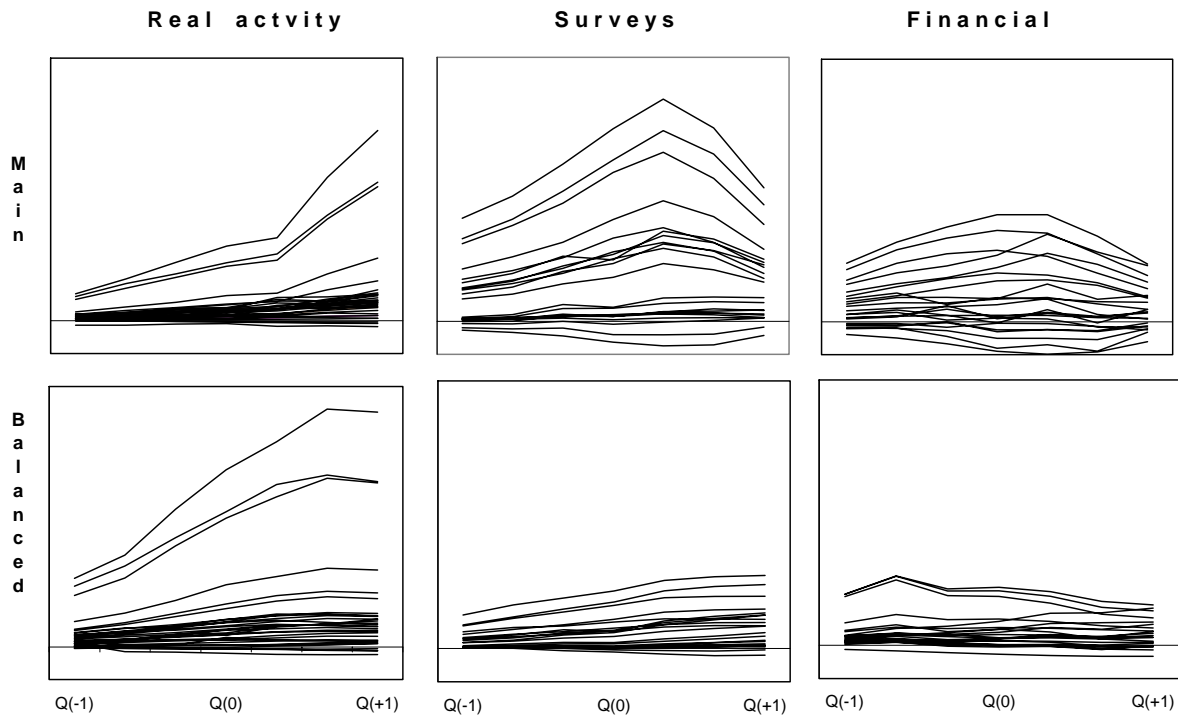
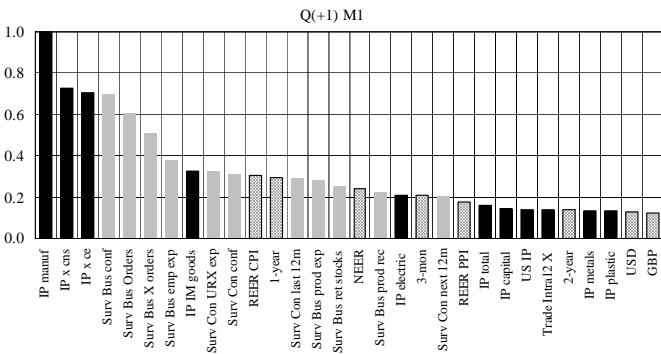
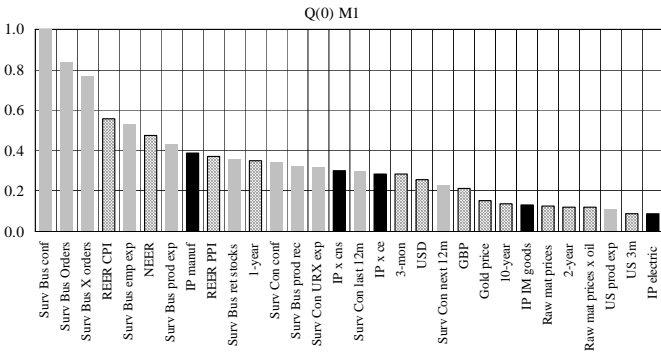
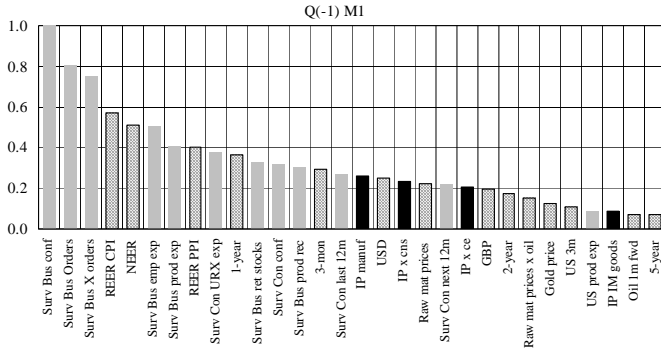


Table 1: Mean absolute contributions (MAC) of data groups  
(1998 Q1 - 2005Q4)

Data	Main				Balanced			
	Fcst	Contributions (%)			Fcst	Contributions (%)		
	$\mathcal{Z}$	$\mathcal{S}$	$\mathcal{F}$	$\mathcal{R}$	$\mathcal{Z}$	$\mathcal{S}$	$\mathcal{F}$	$\mathcal{R}$
Q(-1) M1	0.158	60 %	57 %	14 %	0.135	34 %	46 %	46 %
Q(-1) M2	0.183	58 %	57 %	15 %	0.163	32 %	46 %	44 %
Q(-1) M3	0.196	61 %	56 %	16 %	0.192	28 %	42 %	49 %
Q(0) M1	0.227	62 %	50 %	16 %	0.188	34 %	40 %	48 %
Q(0) M2	0.245	63 %	42 %	17 %	0.199	35 %	35 %	47 %
Q(0) M3	0.230	61 %	37 %	25 %	0.206	35 %	29 %	52 %
Q(+1) M1	0.210	53 %	35 %	32 %	0.200	37 %	29 %	50 %

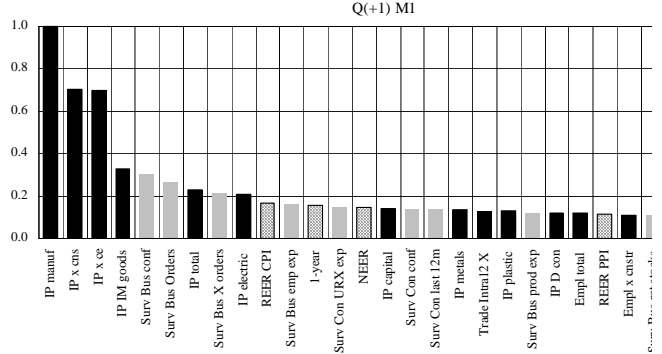
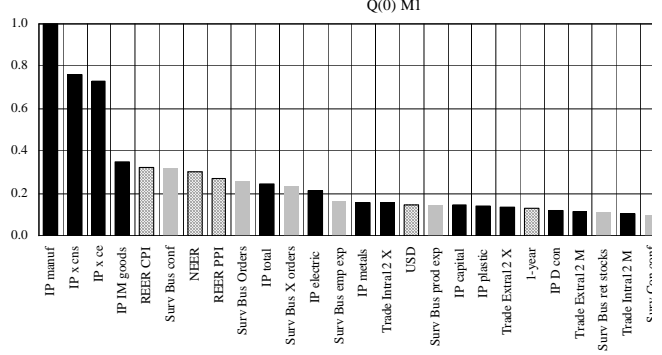
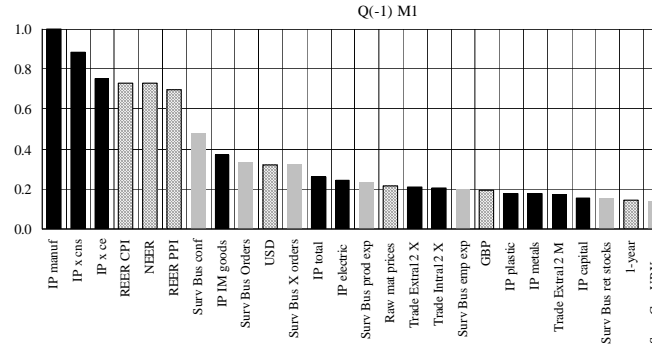
**Chart A.1: Absolute cumulative weights**

Main data



**Chart A.3: Absolute cumulative weights**

Balanced data



**Table 2: Filter uncertainty**

(Full-sample parameter estimates)

	Main				Balanced			
	$Z$	$RS$	$RF$	$SF$	$Z$	$RS$	$RF$	$SF$
Q(-1) M1	.178	.185	.182	.178	.176	.180	.176	.178
Q(-1) M2	.160	.172	.166	.160	.158	.165	.158	.160
Q(-1) M3	.137	.152	.145	.138	.135	.143	.135	.138
Q(0) M1	.100	.112	.119	.100	.091	.097	.091	.100
Q(0) M2	.070	.078	.100	.071	.056	.060	.056	.071
Q(0) M3	.037	.042	.068	.043	.021	.023	.023	.043
Q(+1) M1	.029	.033	.043	.042	.020	.023	.022	.042

Table 3: RMSE from recursive forecasts  
(1998 Q1 - 2005Q4)

	AR	Main				Balanced			
		<i>Z</i>	<i>RS</i>	<i>RF</i>	<i>SF</i>	<i>Z</i>	<i>RS</i>	<i>RF</i>	<i>SF</i>
Q(-1) M1	.38	.33	.37	.32	.33	.33	.35	.32	.33
Q(-1) M2	.35	.32	.36	.31	.32	.31	.33	.31	.32
Q(-1) M3	.35	.28	.33	.29	.28	.28	.30	.29	.28
Q(0) M1	.35	.28	.30	.30	.28	.26	.27	.26	.28
Q(0) M2	.31	.28	.31	.29	.28	.25	.26	.24	.28
Q(0) M3	.31	.25	.28	.27	.27	.24	.25	.24	.27
Q(+1) M1	.31	.24	.25	.24	.27	.23	.24	.23	.27

# Conclusions

Statistics to investigate the role of individual series in forecasts from a DFM

- Based on Kalman filter
- Deals with unbalanced data sets
- No need to re-estimate parameters

Hard & soft data: differences in publication lags matter a lot!

- Surveys are close substitutes to real data - less precise but published earlier
- Financial data provide complementary information

Balanced data give a wrong picture, but many studies have ignored unbalancedness