Breaking the "Iron Rice Bowl:" Evidence of Precautionary Savings from Chinese State-Owned Enterprises Reform¹

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¹The views expressed in this paper are those of the authors and do not necessarily reflect the views of the Federal Reserve Bank of San Francisco.

The elusive quest for precautionary savings?

- Precautionary savings (PS) potentially important for wealth accumulation, esp. for a country with structural changes (China)
- But it's difficult to estimate importance of PS:
 - 1. Hard to identify large and exogenous variations in income uncertainty (Kennickell and Lusardi, 2005, Carroll and Samwick 1998)
 - 2. Hard to separate risks from risk attitude self-selection bias (Fuchs-Schündeln and Schündeln, 2005)
 - 3. Hard to disentangle uncertainty from income expectations (PS or PIH?)
- Estimates of PS range from very small (Dynan 1993; Guiso, et al. 1992) to very large (Carroll-Samwick 1998; Gourinchas-Parker 2002)

Contributions

- 1. Identify income uncertainty using SOE reform as a natural experiment: massive layoffs hit SOEs but not GOV
- 2. Correct self-selection bias related to occupational choice: focus on government-assigned jobs
- 3. Disentangle PS from PIH effects: use information on household income expectations

Main finding: PS accounts for 30% of wealth accumulation of urban SOE workers from 1995 to 2002 (about 6 months of annual income)

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Why is SOE reform a good experiment to use?

- It was big: about 27 million SOE workers were laid off between 1997 and 2002 (China Labor Statistics Yearbook 2003)
- It was largely exogenous and unexpected to individual workers

- It created significant cross-sectional variations of job uncertainty
 - Treatment (SOE): unemployment risk ↑
 - Control (GOV): iron rice bowl kept

Empirical strategy

 Build on models of precautionary savings (Lusardi, 1998; Carroll, Dynan, and Krane, 2003):

$$\frac{W_i}{P_i} = \beta_0 + \frac{\beta_1 SOE_i}{P_i} + \beta_2 RISK_i + \beta_3 \log(P_i) + \beta'_4 Z_i + v_i$$

 Key coefficient β₁: effects of job uncertainty specific to SOE workers

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- Estimate model separately before and after SOE reform
- Identification: diff-in-diff
 - Precautionary savings: $\beta_1^{after} \beta_1^{before} > 0$

The SOE Reform

"Cradle-to-grave" socialism under central planning regime:

- SOE workers and government employees enjoyed similar job security and benefits
- Jobs in both sectors were mostly assigned by government
- Guaranteed employment and pension; near-free housing, education, and health care

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 Starting in late 1990s, many loss-making SOEs were shut down or privatized

From 1997 to 2002, over 27m SOE workers were laid off
Massive layoffs
Who were laid off?

During same period, GOV workers kept the iron rice bowl

 Among individuals who experienced layoffs prior to 2002, 58% worked in SOEs vs 2.3% in government

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The Data

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Chinese Household Income Project surveys (CHIP)

- Conducted by Chinese Academy of Social Science and National Bureau of Statistics (NBS) in 1988, 1995, 2002, and 2007
- Nationally representative and covering 15,000 to 20,000 households in more than 10 provinces
- Focus on CHIP surveys in 1995 and 2002: before and after the SOE reform
 - Focus on prime-aged workers (25-55 years old) in SOE and GOV

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Summary statistics: 1995 vs. 2002

| Variable | 1995 | | 2002 | |
|----------------------|------|-------|------|--------------|
| | Obs. | Mean | Obs. | Mean |
| Financial wealth | 4390 | 10042 | 3027 | 32826 |
| Annual income | 4390 | 7034 | 3027 | 12985 |
| SOE | 4390 | 67.8% | 3027 | 56.2% |
| CV×100 | 4390 | 2.61 | 3027 | 2.9 |
| Male | 4390 | 63.4% | 3027 | 68.8% |
| Health Care | | | | |
| Own payment | 4390 | 9.9% | 3027 | 23.1% |
| Public health care | 4390 | 71.3% | 3027 | 35% |
| Health insurance | 4390 | 8.8% | 3027 | 41.9% |
| Home ownership rate | 4390 | 42% | 3027 | 80.4% |
| Job assigned by Gov. | 4375 | 82.9% | 3018 | 71.9% |

Summary statistics: GOV vs. SOE

| | | 1995 | | | 20 | 02 | |
|-----|------------------|------|-------|-------|------|-------|-------|
| | Variable | Obs. | Mean | SD | Obs. | Mean | SD |
| GOV | Financial wealth | 1414 | 10457 | 10205 | 1325 | 34677 | 32351 |
| | Annual income | 1414 | 7545 | 3214 | 1325 | 14752 | 6698 |
| | W/P | 1414 | 1.376 | 1.386 | 1325 | 2.559 | 2.360 |
| | Non homeowners | 1413 | 0.546 | 0.498 | 1325 | 0.165 | 0.372 |
| | Job assigned | 1409 | 0.893 | 0.309 | 1319 | 0.757 | 0.429 |
| | Exp. income loss | N.A | N.A | N.A | 1321 | 0.114 | 0.318 |
| SOE | Financial wealth | 2976 | 9845 | 10141 | 1702 | 31386 | 31910 |
| | Annual income | 2976 | 6791 | 3385 | 1702 | 11610 | 6294 |
| | W/P | 2976 | 1.382 | 1.448 | 1702 | 2.703 | 2.906 |
| | Non homeowners | 2977 | 0.597 | 0.491 | 1702 | 0.220 | 0.414 |
| | Job assigned | 2966 | 0.798 | 0.401 | 1699 | 0.689 | 0.463 |
| | Exp. income loss | N.A | N.A. | N.A. | 1699 | 0.238 | 0.426 |

Source: CHIP

Empirical Results

Baseline estimation results

| Dep. variable: | 1995 | | 2002 | |
|-------------------------------|--------------------------------|--------------------------------|--------------------------------|--------------------------------|
| W/P | Full sample | Assigned jobs | Full sample | Assigned jobs |
| SOE | 0.039 | 0.090 | 0.327* | 0.723** |
| CV×100 | (0.114) 0.136*** (0.038) | (0.117) 0.111*** (0.040) | (0.221) 0.091*** (0.028) | (0.298) 0.124*** (0.045) |
| Controls Chow-test for SOE | Υ | Υ | Υ | Υ |
| (p-value) | | | 0.247 | 0.048 |
| Log-Likélihood | -8875.88 | -7167.03 | -8240.22 | -5803.38 |
| Sample size | 4390 | 3627 | 3027 | 2170 |

Controls: age, gender, occupation, skills, health care access, marriage, children, #boys, HH size, homeownership, and industry/province dummies. Estimation details

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- All else equal, SOE workers saved slightly more than GOV workers in 1995 (β₁ = 0.039), but difference insignificant
- SOE workers saved significantly more than GOV workers in 2002 ($\beta_1 = 0.327$)

△β₁ identifies diff in W/P due to SOE reform (0.327-0.039=0.288, or 3 months of income)

Self-selection bias (SSB)

- Self selection: occupational choices may be correlated with risk preferences
- Self-selection causes significant downward bias in estimating PS (Fuchs-Schündeln and Schündeln 2005)
- To mitigate SSB, we focus on sample with government-assigned jobs
 - Most jobs in our sample were assigned by government (83% in 1995, 72% in 2002)
 - \blacktriangleright Gov't has final say in job assignments \rightarrow mitigating correlation between occupational choice and worker preferences

Identifying PS controlling for self-selection bias

| Dep. variable: | 1995 | | 2002 | |
|-------------------|-----------------------|------------------------|------------------------|-----------------------|
| W/P | Full sample | Assigned jobs | Full sample | Assigned jobs |
| SOE | 0.039 | 0.090 | 0.327* | 0.723** |
| | (0.114) | (0.117) | (0.221) | (0.298) |
| CV×100 | Ò.136* [*] * | Ò.111* [*] ** | Ò.091* [*] ** | Ò.124* [*] * |
| | (0.038) | (0.040) | (0.028) | (0.045) |
| Controls | Y | Y | Y | Y |
| Chow-test for SOF | • | • | • | |
| (p-value) | | | 0.247 | 0.048 |
| Log-Likelihood | -8875.88 | -7167.03 | -8240.22 | -5803.38 |
| Sample size | 4390 | 3627 | 3027 | 2170 |

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Importance of self-selection bias

- ► No control for self-selection bias: PS = 0.327 - 0.039 = 0.288
- Control for self-selection bias: PS = 0.723 0.090 = 0.633
- Without controlling SSB, PS due to SOE reform would be under-estimated by 0.633-0.288=0.345 (or 4 months of permanent income)—a downward bias of about half of PS
 - Magnitude similar to Fuchs-Schündeln and Schündeln (2005)

- Reform might affect SOE workers' expectations of future income levels
- Lower expected future income may also raise current saving, but such saving reflects wealth effects (or PIH effects): different from PS

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Current estimation mixes PS and PIH effects

How to disentangle PS from PIH?

- 2002 CHIP survey reported households' expected income for next five years: up, down, or unchanged (not reported in 1995 survey)
- We focus on households who expect non-declines in income
- Estimates of PS likely a lower-bound:
 - HH who expected income to fall excluded from sample, but they likely have higher unemployment risks
 - HH who expected income to rise included in sample, but they likely save less

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Regression for 2002 sample controlling for PIH effects

Expected future income

| Dep. variable: W/P | Decline | Non-decline |
|------------------------------|---------|-------------|
| SOE | 1.257** | 0.603** |
| | (0.531) | (0.305) |
| CV×100 | 0.120** | 0.123*** |
| | (0.061) | (0.046) |
| Controls | Y | Υ |
| p-value of Chow test for SOE | 0.032 | 0.116 |
| Ν | 1284 | 1876 |

Note: Sample restricted to government assigned jobs (to control for self-selections).

Quantify Precautionary Savings

- With SSB and PIH both controlled, PS = 0.603-0.09=0.513 (6 months of income)
- Steps to calculate importance of PS:
 - 1. Calculate mean predicted wealth holdings of SOE HH from estimated model: \hat{W}_t^{soe}
 - 2. Calculate *counterfactual* wealth holdings by SOE HH had they faced same job risks as in GOV (by setting SOE = 0): \tilde{W}_t^{soe}
 - 3. Compute magnitude of precautionary savings due to SOE reform

$$W^{ps} \equiv (\hat{W}_{2002}^{soe} - \tilde{W}_{2002}^{soe}) - (\hat{W}_{1995}^{soe} - \tilde{W}_{1995}^{soe})$$

 Contributions of PS to SOE HH wealth accumulation: 30% (likely lower bound)

$$\frac{W^{ps}}{\hat{W}_{2002}^{soe} - \hat{W}_{1995}^{soe}} = 0.303 \quad (s.e. = 0.166)$$

Robustness

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Robustness

Worker composition effects

- survival bias
- voluntary quits
- Other robustness checks:
 - Excluding zero-wealth observations
 - Conventional risk measures
 - Alternative wealth measures
 - Pension effects

For all experiments, we control for self-selection and PIH effects

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- Workers who survived massive layoffs might be different from those before reform • Who were laid off?
- We estimate prob of layoffs for SOE workers using 2002 sample, expanded to include those who had layoff experience

$$\Pr(\mathsf{layoff}_i = 1 \mid Z_i) = \Phi(Z_i \delta + \varepsilon_i)$$

- Then impute prob of layoff for SOE workers in 1995 sample
- Keep only workers in the 1995 sample who are likely to survive reform (with prob of layoff below some threshold)

Voluntary quits

- Some workers quit from SOE for private-sector jobs (quit rate in 2002=1.88%)
- If more risk-averse workers remained in SOE, estimated PS could be biased upward
- To control for effects of quits:
 - 1. Expand 2002 sample to include those who had quit from SOEs to estimate probability of quitting using the Probit model

$$\Pr(\operatorname{quit}_i = 1 \mid Z_i) = \Phi(Z_i \delta + \varepsilon_i)$$

2. Impute probability of quit for SOE workers in 1995 sample; restrict sample to non-quitting workers to make SOE sample comparable between 1995 and 2002

| A. Controlling for survival biases | | | | |
|------------------------------------|------------|-------------|-----------|---------|
| Dep. variable | 1995 sur | vival thres | shold | |
| W/P | 100% | 90% | 80% | 70% |
| SOE | 0.090 | 0.122 | 0.192 | 0.195 |
| | (0.117) | (0.122) | (0.131) | (0.133) |
| Controls | yes | yes | yes | yes |
| Sample size | 3627 | 3415 | 3198 | 2971 |
| B. C | ontrolling | for volunt | ary quits | |
| Dep. variable | 1995 nor | n-quit thre | shold | |
| W/P | 100% | 98% | 96% | 94% |
| SOE | 0.090 | 0.119 | 0.066 | 0.076 |
| | (0.117) | (0.125) | (0.143) | (0.151) |
| Controls | yes | yes | yes | yes |
| Sample size | 3627 | 3582 | 3532 | 3435 |

Other robustness checks

| Case | 1995 | 2002 | Contributions of precautionary savings |
|----------------------|---------|---------|--|
| A. Eliminating zero | 0.100 | 0.467* | 21.8% |
| wealth | (0.104) | (0.268) | (0.133) |
| B. Conventional risk | 0.083 | 0.713** | 37.3% |
| measures | (0.117) | (0.346) | (0.197) |
| C. Very liquid | 0.062 | 0.439* | 33.6% |
| assets | (0.114) | (0.248) | (0.218) |
| D. Non-housing | 0.210 | 0.632* | 29.5% |
| non-business wealth | (0.159) | (0.355) | (0.210) |
| E. Pension effects | 0.09 | 0.580** | 29.4% |
| | (0.117) | (0.307) | (0.172) |

All estimation results here have controlled for self-selection, PIH, and pension effects.

Further Evidence

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Younger households have stronger precautionary saving motive (Gourinchas-Parker, 2002)

| | | 2002 | |
|-------------------|----------|---------|-------------|
| Dep variable: W/P | 25-45 | 46-55 | Full sample |
| SOE | 0.857** | 0.193 | 0.603** |
| | (0.414) | (0.932) | (0.305) |
| CV×100 | 0.145*** | 0.104 | 0.123*** |
| | (0.049) | (0.130) | (0.046) |
| Controls | yes | yes | |
| Sample size | 1087 | 789 | 1876 |

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PS stronger for workers in smaller SOEs

- SOE reform featured "Grasp the large, let go of the small" (Hsieh and Song, 2013)
- Workers in smaller SOEs face higher layoff risks

| Dep. variable: W/P | 1995 | 2002 |
|-----------------------|---------|---------------------|
| CSOE | 0.0001 | 0.088 |
| | (0.146) | (0.294) |
| LSOE | Ò.160 | 1.082* [*] |
| | (0.180) | (0.425) |
| Controls | yes | yes |
| Sample size | 3627 | 1876 |

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Conclusion

- We use the Chinese SOE reform as a natural experiment to identify the existence and importance of precautionary savings
- Our identification of PS takes into account self-selection bias and PIH effects on savings
- We estimate that precautionary savings triggered by SOE reform account for about 30% of the increase in Chinese urban SOE household savings from 1995 to 2002

| Year | SOE layoffs (mi | llion) Effective Urban U (%) | |
|-------|-----------------|------------------------------|---|
| 1997* | 6.92 | 7.7 | - |
| 1998 | 5.62 | 8.5 | |
| 1999 | 6.19 | 9.0 | |
| 2000 | 4.46 | 10.8 | |
| 2001 | 2.34 | 10.8 | |
| 2002 | 1.62 | 11.1 | |
| Total | 27.15 | | _ |

Source: China Labor Statistical Yearbook 2003; Cai, Park, and Zhao (2008); Giles, Park, and Zhang (2005)

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Who Were Laid Off?

| | Never laid off | Experienced layoffs |
|--|----------------------------|-----------------------------------|
| No. of observations | 5770 | 1159 |
| | | |
| Demographics | | |
| Male (%) | 56.8 | 38.7 |
| Education (in years) | 11.4 | 9.96 |
| Not generally healthy (%) | 3.8 | 8.4 |
| | | |
| Ownership (%) | | |
| Central SOEs | 36.8 | 12.1 |
| Local SOEs | 40.9 | 47.8 |
| Urban collective | 9.9 | 31.1 |
| | | |
| Occupation (%) | | |
| Professional/téchnical | 23.2 | 9.5 |
| Administrative/clerical | 31.9 | 13.0 |
| Industrial | 33.0 | 59.1 |
| Commercial and Services | 10.0 | 16.5 |
| Source: 1999 CASS Survey, from Appleton, K | night. Song and Xia (2002) | • Back to Size • Back to Survivor |
| → Back to Reform | 0 ·, · · 0 ····· (-···) | |

Case Study: Lay-off in Fushun, Liaoning

- Fushun is one of state-owned heavy industrial bases in "rust belt" of China
- Before 2000, 91% of workers employed either by SOEs or collective-owned enterprises (COEs)
- In 2000, 42% of SOE and COE workers were laid off, the highest in Liaoning province
- Layoff concentrated in coal, textiles, light industry, electronics, machinery and chemicals
 - 71000 workers in COEs in the coal sector, 35000 or 49.7% of workers were classified as "xia gang" ("left job post")
- Lots of laid-off workers barely got any compensation from firms, but still remained ties with them
- Main avenue for laid-off workers to find new jobs was through re-employment centers sponsored by the local government.
 But re-employment rate was low

Dependent variable: W/P

- Financial wealth (W): checking accounts, saving accounts, CDs, stocks, bonds, and other business assets (Item 401 in CHIP)
 - 1. Financial wealth is not easily affected by high-frequency income fluctuations (unlike flow of saving) \rightarrow mitigates measurement errors
 - 2. It's liquid: useful to safeguard against uncertainty (Carroll and Samwick, 1998)
- Measurement of permanent income (*P*):
 - Constructed using same approach as in Fuchs-Schündeln and Schündeln (2005)

W/P captures cumulative saving



Independent variables

- ► *SOE*: dummy (1 for SOE workers and 0 for GOV)
- RISK: measured by coefficient of variation (CV) of log real income in past years
- ► *P*: permanent income
- Z: demographics (age, gender, HH size, occupation, home ownership, health care, child information, # boys, industry/province, ...)

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Wealth measures

CHIP data

| Α. | Financial wealth | | |
|----|---|--|--|
| | 1. Checking account balances | | |
| | 2. Saving account balances | | |
| | 3. Stocks | | |
| | 4. Bonds | | |
| | 5. Loans to others | | |
| | 6. Own funds for family business | | |
| | 7. Other business assets (excluding stocks and bonds) | | |
| | 8. Housing fund | | |
| | 9. Value of commercial insurance | | |
| | 10. Estimated present market value of collections | | |
| В. | Estimated value of durable goods | | |
| С. | Estimated value of farms and businesses | | |
| D. | Estimated value of houses owned | | |
| E. | Estimated value of other family assets | | |
| F. | Total household debt | | |

Wealth measures:

- Financial wealth: A
- Very liquid assets: A1+A2+A3+A4+A5
- Financial net worth: A-F
- Nonhousing, nonbusiness wealth: A+B+E-F

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Total net worth: A+B+C+D+E-F

Benchmark estimation details • Back

| Dep. variable: | 1995 | | 2002 | |
|------------------------------|--------------------|--------------------|--------------------|-----------------------|
| W/P | (i) | (ii) | (iii) | (iv) |
| SOE | 0.090 | 0.039 | 0.723** | 0.327* |
| | (0.117) | (0.114) | (0.298) | (0.221) |
| CV×100 | 0.111*** | 0.136*** | 0.124*** | 0.091*** |
| | (0.040) | (0.038) | (0.045) | (0.028) |
| log(permanent income) | 0.759 | 1.225 | 4.512*** | 3.533*** |
| | (1.028) | (0.900) | (1.497) | (0.992) |
| Age | ò.020 ′ | -0.020 | 0.028 [′] | 0.240* |
| | (0.052) | (0.050) | (0.150) | (0.125) |
| Age squared(*100) | -0.030 | 0.019 | -0.039 | -0.274* |
| , | (0.059) | (0.059) | (0.175) | (0.147) |
| Male | -0.362*** | -0.463*** | -1.180*** | -1.176*** |
| | (0.102) | (0.094) | (0.202) | (0.148) |
| Professional | 0.102 [′] | ò.031 | 4.776*** | ò.370 [′] |
| | (0.212) | (0.200) | (1.648) | (0.787) |
| Director | 0.295 | 0.185 | 4.780*** | 0.183 |
| | (0.214) | (0.208) | (1.636) | (0.800) |
| Skilled worker | 0.042 | 0.004 | 4.993*** | 0.341 |
| | (0.182) | (0.168) | (1.661) | (0.762) |
| Unskilled worker | -0.031 | 0.039 | 6.093*** | 0.981 |
| | (0.201) | (0.179) | (1.770) | (0.767) |
| Public med service | Ò.047 | 0.036 [′] | -1.228** | -0.978*** |
| | (0.192) | (0.166) | (0.501) | (0.362) |
| Public med insurance | 0.031 | 0.102 | -0.908** | -0.755** |
| | (0.166) | (0.150) | (0.434) | (0.318) |
| Married | 0.520*** | 0.488*** | 0.637 | 0.406 |
| | (0.192) | (0.161) | (0.429) | (0.363) |
| Age of children (mean) | ò.008 ´ | ò.005 ´ | ò.004 | -0.000 |
| | (0.006) | (0.006) | (0.013) | (0.010) |
| Num. of boys | 0.044 | 0.022 | -0.253* | -0.198* |
| - | (0.048) | (0.045) | (0.145) | (0.118) |
| Num. of children at school | -0.086 | -0.035 | -0.317* | -0.363*** |
| | (0.066) | (0.063) | (0.176) | (0.140) |
| Household size | -0.037 | -0.008 | ò.279 ′ | 0.357* ^{***} |
| | (0.051) | (0.048) | (0.171) | (0.136) |
| No house owned | 0.080 | 0.138 | -0.244 | -0.221 |
| | (0.101) | (0.097) | (0.264) | (0.228) |
| No house owned × SOE | -0.114 | -0.106 | 0.356 | 0.300 |
| | (0.109) | (0.104) | (0.376) | (0.300) |
| Industry & Province dummies | yes | yes | yes | yes |
| Log-Likelihood | -7167.03 | -8875.88 | -5803.38 | -8240.22 |
| p-value of Chow test for SOE | | | 0.048 | 0.247 |
| Number of observations | 3627 | 4390 | 2170 | 3027 |